

**Problem 4.7.8 Solution**

The joint PDF of  $X$  and  $Y$  is

$$f_{X,Y}(x,y) = \begin{cases} (x+y)/3 & 0 \leq x \leq 1, 0 \leq y \leq 2 \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

Before calculating moments, we first find the marginal PDFs of  $X$  and  $Y$ . For  $0 \leq x \leq 1$ ,

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy = \int_0^2 \frac{x+y}{3} dy = \frac{xy}{3} + \frac{y^2}{6} \Big|_{y=0}^{y=2} = \frac{2x+2}{3} \quad (2)$$

For  $0 \leq y \leq 2$ ,

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx = \int_0^1 \left(\frac{x}{3} + \frac{y}{3}\right) dx = \frac{x^2}{6} + \frac{xy}{3} \Big|_{x=0}^{x=1} = \frac{2y+1}{6} \quad (3)$$

Complete expressions for the marginal PDFs are

$$f_X(x) = \begin{cases} \frac{2x+2}{3} & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases} \quad f_Y(y) = \begin{cases} \frac{2y+1}{6} & 0 \leq y \leq 2 \\ 0 & \text{otherwise} \end{cases} \quad (4)$$

(a) The expected value of  $X$  is

$$E[X] = \int_{-\infty}^{\infty} xf_X(x) dx = \int_0^1 x \frac{2x+2}{3} dx = \frac{2x^3}{9} + \frac{x^2}{3} \Big|_0^1 = \frac{5}{9} \quad (5)$$

The second moment of  $X$  is

$$E[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_0^1 x^2 \frac{2x+2}{3} dx = \frac{x^4}{6} + \frac{2x^3}{9} \Big|_0^1 = \frac{7}{18} \quad (6)$$

The variance of  $X$  is  $\text{Var}[X] = E[X^2] - (E[X])^2 = 7/18 - (5/9)^2 = 13/162$ .

(b) The expected value of  $Y$  is

$$E[Y] = \int_{-\infty}^{\infty} yf_Y(y) dy = \int_0^2 y \frac{2y+1}{6} dy = \frac{y^2}{12} + \frac{y^3}{9} \Big|_0^2 = \frac{11}{9} \quad (7)$$

The second moment of  $Y$  is

$$E[Y^2] = \int_{-\infty}^{\infty} y^2 f_Y(y) dy = \int_0^2 y^2 \frac{2y+1}{6} dy = \frac{y^3}{18} + \frac{y^4}{12} \Big|_0^2 = \frac{16}{9} \quad (8)$$

The variance of  $Y$  is  $\text{Var}[Y] = E[Y^2] - (E[Y])^2 = 23/81$ .

(c) The correlation of  $X$  and  $Y$  is

$$E[XY] = \iint xyf_{X,Y}(x,y) dx dy \quad (9)$$

$$= \int_0^1 \int_0^2 xy \left(\frac{x+y}{3}\right) dy dx \quad (10)$$

$$= \int_0^1 \left(\frac{x^2 y^2}{6} + \frac{xy^3}{9} \Big|_{y=0}^{y=2}\right) dx \quad (11)$$

$$= \int_0^1 \left(\frac{2x^2}{3} + \frac{8x}{9}\right) dx = \frac{2x^3}{9} + \frac{4x^2}{9} \Big|_0^1 = \frac{2}{3} \quad (12)$$

The covariance is  $\text{Cov}[X, Y] = E[XY] - E[X]E[Y] = -1/81$ .

(d) The expected value of  $X$  and  $Y$  is

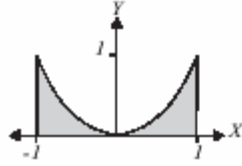
$$E[X+Y] = E[X] + E[Y] = 5/9 + 11/9 = 16/9 \quad (13)$$

(e) By Theorem 4.15,

$$\text{Var}[X+Y] = \text{Var}[X] + \text{Var}[Y] + 2\text{Cov}[X, Y] = \frac{13}{162} + \frac{23}{81} - \frac{2}{81} = \frac{55}{162} \quad (14)$$

**Problem 4.7.10 Solution**

The joint PDF of  $X$  and  $Y$  and the region of nonzero probability are



$$f_{X,Y}(x,y) = \begin{cases} 5x^2/2 & -1 \leq x \leq 1, 0 \leq y \leq x^2 \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

(a) The first moment of  $X$  is

$$E[X] = \int_{-1}^1 \int_0^{x^2} x \frac{5x^2}{2} dy dx = \int_{-1}^1 \frac{5x^5}{2} dx = \left. \frac{5x^6}{12} \right|_{-1}^1 = 0 \quad (2)$$

Since  $E[X] = 0$ , the variance of  $X$  and the second moment are both

$$\text{Var}[X] = E[X^2] = \int_{-1}^1 \int_0^{x^2} x^2 \frac{5x^2}{2} dy dx = \left. \frac{5x^7}{14} \right|_{-1}^1 = \frac{10}{14} \quad (3)$$

(b) The first and second moments of  $Y$  are

$$E[Y] = \int_{-1}^1 \int_0^{x^2} y \frac{5x^2}{2} dy dx = \frac{5}{14} \quad (4)$$

$$E[Y^2] = \int_{-1}^1 \int_0^{x^2} x^2 y^2 \frac{5x^2}{2} dy dx = \frac{5}{26} \quad (5)$$

Therefore,  $Y$  has variance

$$\text{Var}[Y] = \frac{5}{26} - \left( \frac{5}{14} \right)^2 = .0576 \quad (6)$$

(c) Since  $E[X] = 0$ ,  $\text{Cov}[X, Y] = E[XY] - E[X]E[Y] = E[XY]$ . Thus,

$$\text{Cov}[X, Y] = E[XY] = \int_{-1}^1 \int_0^{x^2} xy \frac{5x^2}{2} dy dx = \int_{-1}^1 \frac{5x^7}{4} dx = 0 \quad (7)$$

(d) The expected value of the sum  $X + Y$  is

$$E[X + Y] = E[X] + E[Y] = \frac{5}{14} \quad (8)$$

(e) By Theorem 4.15, the variance of  $X + Y$  is

$$\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y] + 2\text{Cov}[X, Y] = 5/7 + 0.0576 = 0.7719 \quad (9)$$

### Problem 6.2.2 Solution

The joint PDF of  $X$  and  $Y$  is

$$f_{X,Y}(x,y) = \begin{cases} 1 & 0 \leq x, y \leq 1 \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

Proceeding as in Problem 6.2.1, we must first find  $F_W(w)$  by integrating over the square defined by  $0 \leq x, y \leq 1$ . Again we are forced to find  $F_W(w)$  in parts as we did in Problem 6.2.1 resulting in the following integrals for their appropriate regions. For  $0 \leq w \leq 1$ ,

$$F_W(w) = \int_0^w \int_0^{w-x} dx dy = w^2/2 \quad (2)$$

For  $1 \leq w \leq 2$ ,

$$F_W(w) = \int_0^{w-1} \int_0^1 dx dy + \int_{w-1}^1 \int_0^{w-y} dx dy = 2w - 1 - w^2/2 \quad (3)$$

The complete expression for the CDF of  $W$  is

$$F_W(w) = \begin{cases} 0 & w < 0 \\ w^2/2 & 0 \leq w \leq 1 \\ 2w - 1 - w^2/2 & 1 \leq w \leq 2 \\ 1 & \text{otherwise} \end{cases} \quad (4)$$

With the CDF, we can find  $f_W(w)$  by differentiating with respect to  $w$ .

$$f_W(w) = \begin{cases} w & 0 \leq w \leq 1 \\ 2 - w & 1 \leq w \leq 2 \\ 0 & \text{otherwise} \end{cases} \quad (5)$$

### Problem 6.2.3 Solution

By using Theorem 6.5, we can find the PDF of  $W = X + Y$  by convolving the two exponential distributions. For  $\mu \neq \lambda$ ,

$$f_W(w) = \int_{-\infty}^{\infty} f_X(x)f_Y(w-x) dx \quad (1)$$

$$= \int_0^w \lambda e^{-\lambda x} \mu e^{-\mu(w-x)} dx \quad (2)$$

$$= \lambda \mu e^{-\mu w} \int_0^w e^{-(\lambda-\mu)x} dx \quad (3)$$

$$= \begin{cases} \frac{\lambda \mu}{\lambda - \mu} (e^{-\mu w} - e^{-\lambda w}) & w \geq 0 \\ 0 & \text{otherwise} \end{cases} \quad (4)$$

When  $\mu = \lambda$ , the previous derivation is invalid because of the denominator term  $\lambda - \mu$ . For  $\mu = \lambda$ , we have

$$f_W(w) = \int_{-\infty}^{\infty} f_X(x)f_Y(w-x) dx \quad (5)$$

$$= \int_0^w \lambda e^{-\lambda x} \lambda e^{-\lambda(w-x)} dx \quad (6)$$

$$= \lambda^2 e^{-\lambda w} \int_0^w dx \quad (7)$$

$$= \begin{cases} \lambda^2 w e^{-\lambda w} & w \geq 0 \\ 0 & \text{otherwise} \end{cases} \quad (8)$$

Note that when  $\mu = \lambda$ ,  $W$  is the sum of two iid exponential random variables and has a second order Erlang PDF.

### Problem 10.2.1 Solution

- In Example 10.3, the daily noontime temperature at Newark Airport is a discrete time, continuous value random process. However, if the temperature is recorded only in units of one degree, then the process would be discrete value.
- In Example 10.4, the the number of active telephone calls is discrete time and discrete value.
- The dice rolling experiment of Example 10.5 yields a discrete time, discrete value random process.
- The QPSK system of Example 10.6 is a continuous time and continuous value random process.

### Problem 10.2.4 Solution

The statement is *false*. As a counterexample, consider the rectified cosine waveform  $X(t) = R|\cos 2\pi ft|$  of Example 10.9. When  $t = \pi/2$ , then  $\cos 2\pi ft = 0$  so that  $X(\pi/2) = 0$ . Hence  $X(\pi/2)$  has PDF

$$f_{X(\pi/2)}(x) = \delta(x) \quad (1)$$

That is,  $X(\pi/2)$  is a discrete random variable.

**B.** From the textbook, two random variables X and Y are orthogonal if their correlation,  $r_{X,Y}=0$ .

Thus, we want  $r_{V,W} = 0$ ;

$$r_{V,W} = E[VW]$$

$$r_{V,W} = E[(X + aY)(X - aY)]$$

$$r_{V,W} = E[X^2 + aXY - aXY - a^2Y^2]$$

$$r_{V,W} = E[X^2] - a^2E[Y^2]$$

Now we set  $r_{V,W} = 0$

$$E[X^2] - a^2E[Y^2] = 0$$

And solve for a:

$$a = \pm \sqrt{\left(\frac{E[X^2]}{E[Y^2]}\right)}$$